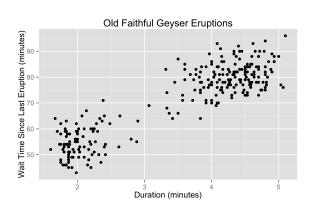
#### K-Means and Gaussian Mixture Models

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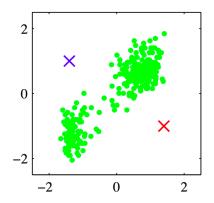
# Example: Old Faithful Geyser



- Looks like two clusters.
- How to find these clusters algorithmically?

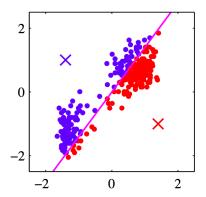
# k-Means: By Example

- Standardize the data.
- Choose two cluster centers.

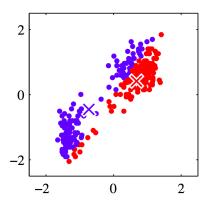


From Bishop's Pattern recognition and machine learning, Figure 9.1(a).

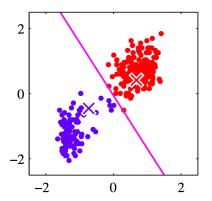
Assign each point to closest center.



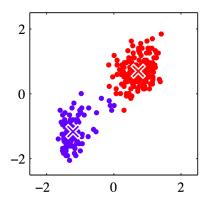
• Compute new class centers.



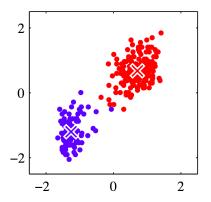
Assign points to closest center.



Compute cluster centers.



• Iterate until convergence.



#### k-means: formalization

- Dataset  $\mathcal{D} = \{x_1, \dots, x_n\} \in \mathbb{R}^d$
- Goal (version 1): Partition data into k clusters.
- Goal (version 2): Partition  $\mathbb{R}^d$  into k regions.
- Let  $\mu_1, \ldots, \mu_k$  denote cluster centers.

#### k-means: formalization

• For each  $x_i$ , use a **one-hot encoding** to designate membership:

$$r_i = (0, 0, \dots, 0, 0, 1, 0, 0) \in \mathbb{R}^k$$

Let

$$r_{ic} = 1(x_i \text{ assigned to cluster } c)$$
.

Then

$$r_i = (r_{i1}, r_{i2}, \ldots, r_{ik}).$$

# k-means: objective function

Find cluster centers and cluster assignments minimizing

$$J(r, \mu) = \sum_{i=1}^{n} \sum_{c=1}^{k} r_{ic} ||x_i - \mu_c||^2.$$

- Is objective function convex?
- What's the domain of *J*?
- $r \in \{0,1\}^{n \times k}$ , which is not a convex set...
- So domain of J is not convex  $\implies$  J is not a convex function
- We should expect local minima.
- Could replace  $\|\cdot\|^2$  with something else:
  - e.g. using  $\|\cdot\|$  (or any distance metric) gives k-medoids.

### k-means algorithm

• For fixed r (cluster assignments), minimizing over  $\mu$  is easy:

$$J(r, \mu) = \sum_{i=1}^{n} \sum_{c=1}^{k} r_{ic} \|x_i - \mu_c\|^2$$

$$= \sum_{c=1}^{k} \sum_{i=1}^{n} r_{ic} \|x_i - \mu_c\|^2$$

$$J_c(\mu_c) = \sum_{\{i \mid x_i \text{ belongs to cluster } c\}} \|x_i - \mu_c\|^2$$

• J<sub>c</sub> is minimized by

$$\mu_c = \text{mean}(\{x_i \mid x_i \text{ belongs to cluster } c\})$$

## k-means algorithm

• For fixed  $\mu$  (cluster centers), minimizing over r is easy:

$$J(r, \mu) = \sum_{i=1}^{n} \sum_{c=1}^{k} r_{ic} ||x_i - \mu_c||^2$$

• For each *i*, exactly one of the following terms is nonzero:

$$r_{i1}||x_i - \mu_1||^2, r_{i2}||x_i - \mu_2||^2, \dots, r_{ik}||x_i - \mu_k||^2$$

Take

$$r_{ic} = 1(c = \underset{j}{\operatorname{arg\,min}} \|x_i - \mu_j\|^2)$$

• That is, assign  $x_i$  to cluster c with minimum distance

$$||x_i - \mu_c||^2$$

# k-means algorithm (summary)

- We will use an alternating minimization algorithm:
  - **1** Choose initial cluster centers  $\mu = (\mu_1, ..., \mu_k)$ .
    - e.g. choose k randomly chosen data points
  - 2 Repeat
    - For given cluster centers, find optimal cluster assignments:

$$r_{ic}^{\mathsf{new}} = 1(c = \operatorname*{arg\,min}_{j} \|x_i - \mu_j\|^2)$$

Q Given cluster assignments, find optimal cluster centers:

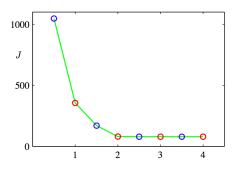
$$\mu_c^{\text{new}} = \underset{m \in \mathbb{R}^d}{\text{arg min}}; \sum_{\{i \mid r_{ic} = 1\}} \|x_i - \mu_c\|^2$$

# k-Means Algorithm: Convergence

- Note: Objective value never increases in an update.
  - (Obvious: worst case, everything stays the same)
- Consider the sequence of objective values:  $J_1, J_2, J_3, \dots$ 
  - monotonically decreasing
  - bounded below by zero
- Therefore, k-Means objective value converges to  $\inf_t J_t$ .
- Reminder: This is convergence to a local minimum.
- Best to repeat k-means several times, with different starting points

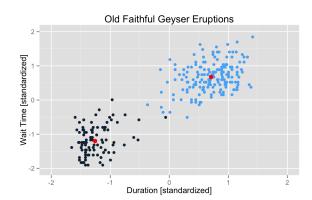
# k-Means: Objective Function Convergence

- Blue circles after "E" step: assigning each point to a cluster
- Red circles after "M" step: recomputing the cluster centers



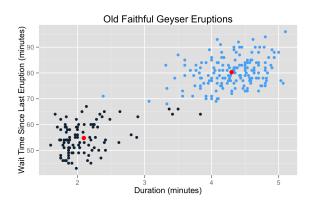
# k-Means Algorithm: Standardizing the data

• With standardizing:



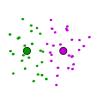
# k-Means Algorithm: Standardizing the data

Without standardizing:

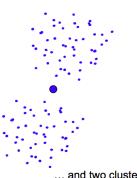


# k-Means: Suboptimal Local Minimum

• The clustering for k = 3 below is a local minimum, but suboptimal:



Would be better to have one cluster here



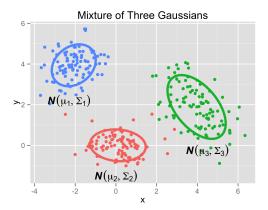
and two clusters here

# Probabilistic Model for Clustering

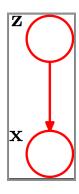
- Let's consider a generative model for the data.
- Suppose
  - $\bullet$  There are k clusters.
  - We have a probability density for each cluster.
- Generate a point as follows
  - ① Choose a random cluster  $z \in \{1, 2, ..., k\}$ .
    - $Z \sim \mathsf{Multi}(\pi_1, \ldots, \pi_k)$ .
  - 2 Choose a point from the distribution for cluster Z.
    - $X \mid Z = z \sim p(x \mid z)$ .

# Gaussian Mixture Model (k = 3)

- **1** Choose  $Z \in \{1, 2, 3\} \sim \text{Multi}\left(\frac{1}{3}, \frac{1}{3}, \frac{1}{3}\right)$ .
- 2 Choose  $X \mid Z = z \sim \mathcal{N}(X \mid \mu_z, \Sigma_z)$ .



#### Gaussian Mixture Model: Joint Distribution



• Factorize joint according to Bayes net:

$$p(x,z) = p(z)p(x \mid z)$$
  
=  $\pi_z \mathcal{N}(x \mid \mu_z, \Sigma_z)$ 

- $\pi_z$  is probability of choosing cluster z.
- $X \mid Z = z$  has distribution  $\mathcal{N}(\mu_z, \Sigma_z)$ .
- z corresponding to x is the true cluster assignment.

#### Latent Variable Model

- Back in reality, we observe X, not (X, Z).
- Cluster assignment Z is called a hidden variable.

#### Definition

A **latent variable model** is a probability model for which certain variables are never observed.

• e.g. The Gaussian mixture model is a latent variable model.

# Model-Based Clustering

- We observe X = x.
- The conditional distribution of the cluster Z given X = x is

$$p(z \mid X = x) = p(x, z)/p(x)$$

- The conditional distribution is a soft assignment to clusters.
- A hard assignment is

$$z^* = \underset{z \in \{1, \dots, k\}}{\operatorname{arg\,min}} \mathbb{P}(Z = z \mid X = x).$$

• So if we have the model, clustering is trival.

# Estimating/Learning the Gaussian Mixture Model

- We'll use the common acronym **GMM**.
- What does it mean to "have" or "know" the GMM?
- It means knowing the parameters

Cluster probabilities: 
$$\pi = (\pi_1, \dots, \pi_k)$$

Cluster means : 
$$\mu = (\mu_1, ..., \mu_k)$$

Cluster covariance matrices: 
$$\Sigma = (\Sigma_1, \dots \Sigma_k)$$

- We have a probability model: let's find the MLE.
- Suppose we have data  $\mathcal{D} = \{x_1, \dots, x_n\}$ .
- We need the model likelihood for  $\mathfrak{D}$ .

### Gaussian Mixture Model: Marginal Distribution

• Since we only observe X, we need the marginal distribution:

$$p(x) = \sum_{z=1}^{k} p(x, z)$$
$$= \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x \mid \mu_{z}, \Sigma_{z})$$

- Note that p(x) is a convex combination of probability densities.
- This is a common form for a probability model...

# Mixture Distributions (or Mixture Models)

#### Definition

A probability density p(x) represents a mixture distribution or mixture model, if we can write it as a convex combination of probability densities. That is,

$$p(x) = \sum_{i=1}^{k} w_i p_i(x),$$

where  $w_i \ge 0$ ,  $\sum_{i=1}^k w_i = 1$ , and each  $p_i$  is a probability density.

- In our Gaussian mixture model, X has a mixture distribution.
- More constructively, let *S* be a set of probability distributions:
  - **1** Choose a distribution randomly from *S*.
  - Sample X from the chosen distribution.
- Then X has a mixture distribution.

# Estimating/Learning the Gaussian Mixture Model

• The model likelihood for  $\mathcal{D} = \{x_1, \dots, x_n\}$  is

$$L(\pi, \mu, \Sigma) = \prod_{i=1}^{n} p(x_i)$$
$$= \prod_{i=1}^{n} \sum_{z=1}^{k} \pi_z \mathcal{N}(x_i \mid \mu_z, \Sigma_z).$$

• As usual, we'll take our objective function to be the log of this:

$$J(\pi, \mu, \Sigma) = \sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}$$

# Properties of the GMM Log-Likelihood

GMM log-likelihood:

$$J(\pi, \mu, \Sigma) = \sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}$$

• Let's compare to the log-likelihood for a single Gaussian:

$$\sum_{i=1}^{n} \log \mathcal{N}(x_i \mid \mu, \Sigma)$$

$$= -\frac{nd}{2} \log (2\pi) - \frac{n}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)' \Sigma^{-1}(x_i - \mu)$$

- For a single Gaussian, the log cancels the exp in the Gaussian density.
  - ⇒ Things simplify a lot.
- For the GMM, the sum inside the log prevents this cancellation.
  - Expression more complicated. No closed form expression for MLE.

# Identifiability Issues for GMM

Suppose we have found parameters

Cluster probabilities:  $\pi = (\pi_1, ..., \pi_k)$ 

Cluster means:  $\mu = (\mu_1, \dots, \mu_k)$ 

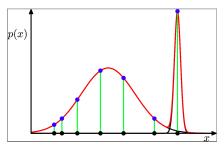
Cluster covariance matrices:  $\Sigma = (\Sigma_1, \dots \Sigma_k)$ 

that are at a local minimum.

- What happens if we shuffle the clusters? e.g. Switch the labels for clusters 1 and 2.
- We'll get the same likelihood. How many such equivalent settings are there?
- Assuming all clusters are distinct, there are k! equivalent solutions.
- Not a problem per se, but something to be aware of.

# Singularities for GMM

Consider the following GMM for 7 data points:



- Let  $\sigma^2$  be the variance of the skinny component.
- What happens to the likelihood as  $\sigma^2 \rightarrow 0$ ?
- In practice, we end up in local minima that do not have this problem.
  - Or keep restarting optimization until we do.
- Bayesian approach or regularization will also solve the problem.

From Bishop's Pattern recognition and machine learning, Figure 9.7.

# Gradient Descent / SGD for GMM

What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

- Can be done but need to be clever about it.
- Each matrix  $\Sigma_1, \ldots, \Sigma_k$  has to be positive semidefinite.
- How to maintain that constraint?
  - Rewrite  $\Sigma_i = M_i M_i^T$ , where  $M_i$  is an unconstrained matrix.
  - Then  $\Sigma_i$  is positive semidefinite.
- But we actually prefer positive definite, to avoid singularities.

# Cholesky Decomposition for SPD Matrices

#### **Theorem**

Every symmetric positive definite matrix  $A \in \mathbb{R}^{d \times d}$  has a unique **Cholesky** decomposition:

$$A = LL^T$$
,

where L a lower triangular matrix with positive diagonal elements.

- A lower triangular matrix has half the number of parameters.
- Symmetric positive definite is better because avoids singularities.
- Requires a non-negativity constraint on diagonal elements.
  - e.g. Use projected SGD method like we did for the Lasso.

#### MLE for Gaussian Model

- Let's start by considering the MLE for the Gaussian model.
- For data  $\mathcal{D} = \{x_1, \dots, x_n\}$ , the log likelihood is given by

$$\sum_{i=1}^{n} \log \mathcal{N}(x_i \mid \mu, \Sigma) = -\frac{nd}{2} \log (2\pi) - \frac{n}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)' \Sigma^{-1}(x_i - \mu).$$

With some calculus, we find that the MLE parameters are

$$\mu_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^{n} x_{i}$$

$$\Sigma_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^{n} (x_{i} - \mu_{\text{MLE}}) (x_{i} - \mu_{\text{MLE}})^{T}$$

- For GMM, If we knew the cluster assignment  $z_i$  for each  $x_i$ ,
  - we could compute the MLEs for each cluster.

# Cluster Responsibilities: Some New Notation

ullet Denote the probability that observed value  $x_i$  comes from cluster j by

$$\gamma_i^j = \mathbb{P}(Z = j \mid X = x_i).$$

- The **responsibility** that cluster j takes for observation  $x_i$ .
- Computationally,

$$\gamma_i^j = \mathbb{P}(Z = j \mid X = x_i). 
= p(Z = j, X = x_i)/p(x) 
= \frac{\pi_j \mathcal{N}(x_i \mid \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i \mid \mu_c, \Sigma_c)}$$

- The vector  $(\gamma_i^1, \dots, \gamma_i^k)$  is exactly the **soft assignment** for  $x_i$ .
- Let  $n_c = \sum_{i=1}^n \gamma_i^c$  be the number of points "soft assigned" to cluster c.

## EM Algorithm for GMM: Overview

- **1** Initialize parameters  $\mu$ ,  $\Sigma$ ,  $\pi$ .
- (a) "E step". Evaluate the responsibilities using current parameters:

$$\gamma_i^j = \frac{\pi_j \mathcal{N}(x_i \mid \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i \mid \mu_c, \Sigma_c)},$$

for i = 1, ..., n and j = 1, ..., k.

3 "M step". Re-estimate the parameters using responsibilities:

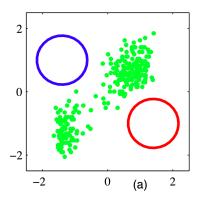
$$\mu_c^{\text{new}} = \frac{1}{n_c} \sum_{i=1}^n \gamma_i^c x_i$$

$$\Sigma_c^{\text{new}} = \frac{1}{n_c} \sum_{i=1}^n \gamma_i^c (x_i - \mu_{\text{MLE}}) (x_i - \mu_{\text{MLE}})^T$$

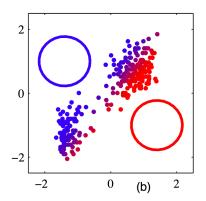
$$\pi_c^{\text{new}} = \frac{n_c}{n},$$

Repeat from Step 2, until log-likelihood converges.

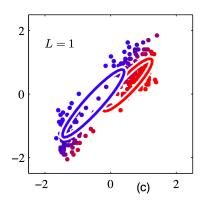
Initialization



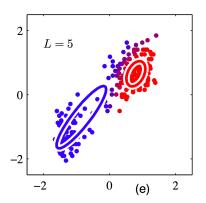
• First soft assignment:



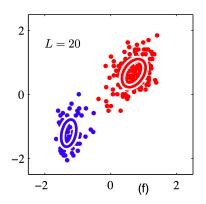
• First soft assignment:



After 5 rounds of EM:



After 20 rounds of EM:



#### Relation to K-Means

- EM for GMM seems a little like k-means.
- In fact, there is a precise correspondence.
- First, fix each cluster covariance matrix to be  $\sigma^2 I$ .
- As we take  $\sigma^2 \to 0$ , the update equations converge to doing *k*-means.
- If you do a quick experiment yourself, you'll find
  - Soft assignments converge to hard assignments.
  - Has to do with the tail behavior (exponential decay) of Gaussian.

## Possible Topics for Next Time

- In last lecture, will give high level view of several topics.
- Possibilities:
  - General EM Algorithm.
  - Bandit problems.
  - LDA / Topic Models
  - Ranking problems.
  - Collaborative Filtering.
  - Generalization bounds.
  - Sequence models (maximum entropy Markov models, HMMs)