

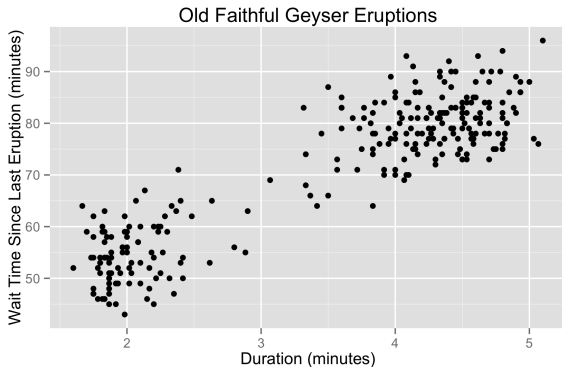
K-Means and Gaussian Mixture Models

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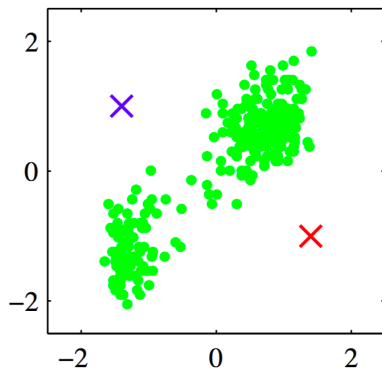
Example: Old Faithful Geyser



- Looks like two clusters.
- How to find these clusters algorithmically?

k-Means: By Example

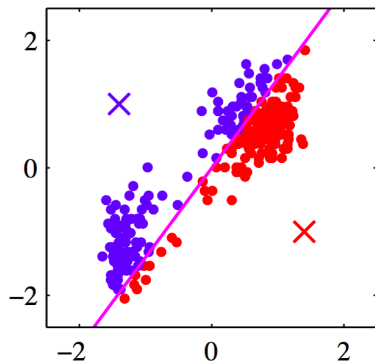
- Standardize the data.
- Choose two cluster centers.



From Bishop's *Pattern recognition and machine learning*, Figure 9.1(a).

k-means: by example

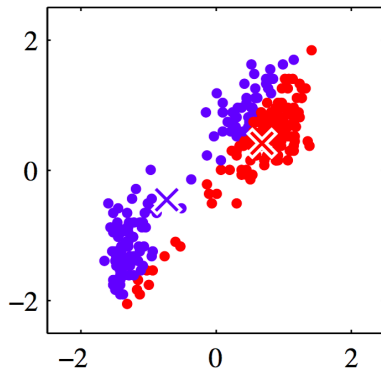
- Assign each point to closest center.



From Bishop's *Pattern recognition and machine learning*, Figure 9.1(b).

k-means: by example

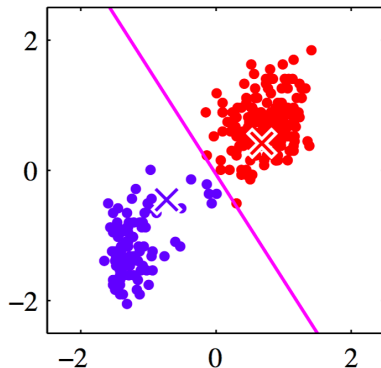
- Compute new class centers.



From Bishop's *Pattern recognition and machine learning*, Figure 9.1(c).

k-means: by example

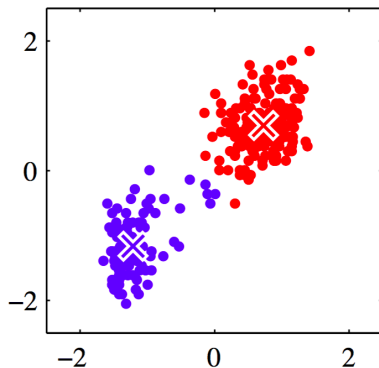
- Assign points to closest center.



From Bishop's *Pattern recognition and machine learning*, Figure 9.1(d).

k-means: by example

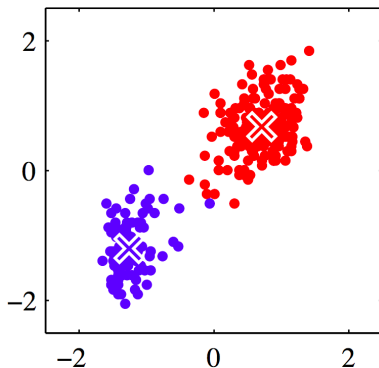
- Compute cluster centers.



From Bishop's *Pattern recognition and machine learning*, Figure 9.1(e).

k-means: by example

- Iterate until convergence.



From Bishop's *Pattern recognition and machine learning*, Figure 9.1(i).

k-means: formalization

- Dataset $\mathcal{D} = \{x_1, \dots, x_n\} \in \mathbf{R}^d$
- Goal (version 1): Partition data into k clusters.
- Goal (version 2): Partition \mathbf{R}^d into k regions.
- Let μ_1, \dots, μ_k denote cluster centers.

k-means: formalization

- For each x_i , use a **one-hot encoding** to designate membership:

$$r_i = (0, 0, \dots, 0, 0, 1, 0, 0) \in \mathbf{R}^k$$

- Let

$$r_{ic} = 1(x_i \text{ assigned to cluster } c).$$

- Then

$$r_i = (r_{i1}, r_{i2}, \dots, r_{ik}).$$

k-means: objective function

- Find cluster centers and cluster assignments minimizing

$$J(r, \mu) = \sum_{i=1}^n \sum_{c=1}^k r_{ic} \|x_i - \mu_c\|^2.$$

- Is objective function convex?
- What's the domain of J ?
- $r \in \{0, 1\}^{n \times k}$, which is not a convex set...
- So domain of J is not convex $\implies J$ is not a convex function
- We should expect local minima.
- Could replace $\|\cdot\|^2$ with something else:
 - e.g. using $\|\cdot\|$ (or any distance metric) gives k -**medoids**.

k-means algorithm

- For fixed r (cluster assignments), minimizing over μ is easy:

$$\begin{aligned}
 J(r, \mu) &= \sum_{i=1}^n \sum_{c=1}^k r_{ic} \|x_i - \mu_c\|^2 \\
 &= \sum_{c=1}^k \underbrace{\sum_{i=1}^n r_{ic} \|x_i - \mu_c\|^2}_{=J_c} \\
 J_c(\mu_c) &= \sum_{\{i|x_i \text{ belongs to cluster } c\}} \|x_i - \mu_c\|^2
 \end{aligned}$$

- J_c is minimized by

$$\mu_c = \text{mean}(\{x_i \mid x_i \text{ belongs to cluster } c\})$$

k-means algorithm

- For fixed μ (cluster centers), minimizing over r is easy:

$$J(r, \mu) = \sum_{i=1}^n \sum_{c=1}^k r_{ic} \|x_i - \mu_c\|^2$$

- For each i , exactly one of the following terms is nonzero:

$$r_{i1} \|x_i - \mu_1\|^2, r_{i2} \|x_i - \mu_2\|^2, \dots, r_{ik} \|x_i - \mu_k\|^2$$

- Take

$$r_{ic} = 1(c = \arg \min_j \|x_i - \mu_j\|^2)$$

- That is, assign x_i to cluster c with minimum distance

$$\|x_i - \mu_c\|^2$$

k-means algorithm (summary)

- We will use an **alternating minimization** algorithm:
 - ① Choose initial cluster centers $\mu = (\mu_1, \dots, \mu_k)$.
 - e.g. choose k randomly chosen data points
 - ② Repeat
 - ① For given cluster centers, find optimal cluster assignments:

$$r_{ic}^{\text{new}} = 1(c = \arg \min_j \|x_i - \mu_j\|^2)$$

- ② Given cluster assignments, find optimal cluster centers:

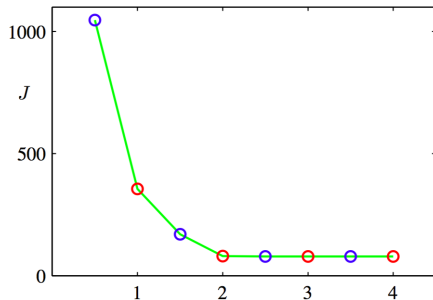
$$\mu_c^{\text{new}} = \arg \min_{m \in \mathbf{R}^d} \sum_{\{i | r_{ic} = 1\}} \|x_i - \mu_c\|^2$$

k-Means Algorithm: Convergence

- Note: Objective value never increases in an update.
 - (Obvious: worst case, everything stays the same)
- Consider the sequence of objective values: J_1, J_2, J_3, \dots
 - monotonically decreasing
 - bounded below by zero
- Therefore, **k-Means objective value converges** to $\inf_t J_t$.
- **Reminder:** This is convergence to a **local** minimum.
- Best to repeat *k*-means several times, with different starting points

k-Means: Objective Function Convergence

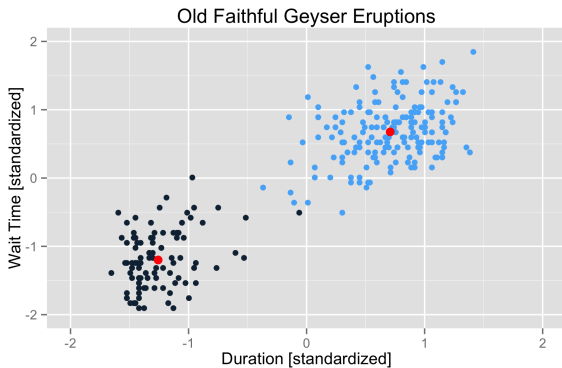
- Blue circles after “E” step: assigning each point to a cluster
- Red circles after “M” step: recomputing the cluster centers



From Bishop's *Pattern recognition and machine learning*, Figure 9.2.

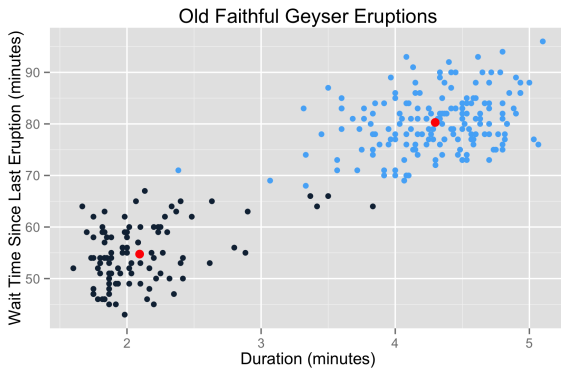
k-Means Algorithm: Standardizing the data

- With standardizing:



k-Means Algorithm: Standardizing the data

- Without standardizing:

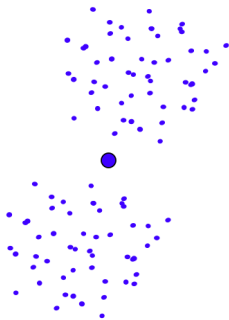


k-Means: Suboptimal Local Minimum

- The clustering for $k = 3$ below is a local minimum, but suboptimal:



Would be better to have
one cluster here



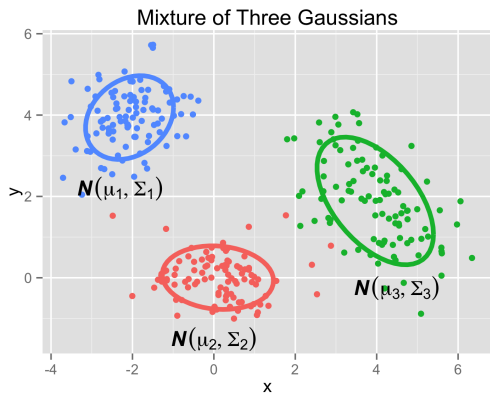
... and two clusters here

Probabilistic Model for Clustering

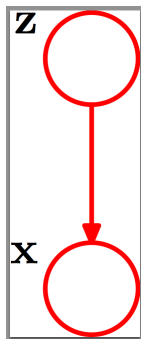
- Let's consider a **generative model** for the data.
- Suppose
 - ① There are k clusters.
 - ② We have a probability density for each cluster.
- Generate a point as follows
 - ① Choose a random cluster $z \in \{1, 2, \dots, k\}$.
 - $Z \sim \text{Multi}(\pi_1, \dots, \pi_k)$.
 - ② Choose a point from the distribution for cluster Z .
 - $X | Z = z \sim p(x | z)$.

Gaussian Mixture Model ($k = 3$)

- 1 Choose $Z \in \{1, 2, 3\} \sim \text{Multi}(\frac{1}{3}, \frac{1}{3}, \frac{1}{3})$.
- 2 Choose $X | Z = z \sim \mathcal{N}(X | \mu_z, \Sigma_z)$.



Gaussian Mixture Model: Joint Distribution



- Factorize joint according to Bayes net:

$$\begin{aligned} p(x, z) &= p(z)p(x | z) \\ &= \pi_z \mathcal{N}(x | \mu_z, \Sigma_z) \end{aligned}$$

- π_z is probability of choosing cluster z .
- $X | Z = z$ has distribution $\mathcal{N}(\mu_z, \Sigma_z)$.
- z corresponding to x is the true cluster assignment.

Latent Variable Model

- Back in reality, we observe X , not (X, Z) .
- Cluster assignment Z is called a **hidden variable**.

Definition

A **latent variable model** is a probability model for which certain variables are never observed.

- e.g. The Gaussian mixture model is a latent variable model.

Model-Based Clustering

- We observe $X = x$.
- The conditional distribution of the cluster Z given $X = x$ is

$$p(z | X = x) = p(x, z) / p(x)$$

- The conditional distribution is a **soft assignment** to clusters.
- A **hard assignment** is

$$z^* = \operatorname{arg\,min}_{z \in \{1, \dots, k\}} \mathbb{P}(Z = z | X = x).$$

- So if we have the model, clustering is trivial.

Estimating/Learning the Gaussian Mixture Model

- We'll use the common acronym **GMM**.
- What does it mean to “have” or “know” the GMM?
- It means knowing the parameters

Cluster probabilities: $\pi = (\pi_1, \dots, \pi_k)$

Cluster means: $\mu = (\mu_1, \dots, \mu_k)$

Cluster covariance matrices: $\Sigma = (\Sigma_1, \dots, \Sigma_k)$

- We have a probability model: let's find the MLE.
- Suppose we have data $\mathcal{D} = \{x_1, \dots, x_n\}$.
- We need the model likelihood for \mathcal{D} .

Gaussian Mixture Model: Marginal Distribution

- Since we only observe X , we need the **marginal distribution**:

$$\begin{aligned} p(x) &= \sum_{z=1}^k p(x, z) \\ &= \sum_{z=1}^k \pi_z \mathcal{N}(x \mid \mu_z, \Sigma_z) \end{aligned}$$

- Note that $p(x)$ is a convex combination of probability densities.
- This is a common form for a probability model...

Mixture Distributions (or Mixture Models)

Definition

A probability density $p(x)$ represents a **mixture distribution** or **mixture model**, if we can write it as a **convex combination** of probability densities. That is,

$$p(x) = \sum_{i=1}^k w_i p_i(x),$$

where $w_i \geq 0$, $\sum_{i=1}^k w_i = 1$, and each p_i is a probability density.

- In our Gaussian mixture model, X has a **mixture distribution**.
- More constructively, let S be a set of probability distributions:
 - 1 Choose a distribution randomly from S .
 - 2 Sample X from the chosen distribution.
- Then X has a mixture distribution.

Estimating/Learning the Gaussian Mixture Model

- The model likelihood for $\mathcal{D} = \{x_1, \dots, x_n\}$ is

$$\begin{aligned} L(\pi, \mu, \Sigma) &= \prod_{i=1}^n p(x_i) \\ &= \prod_{i=1}^n \sum_{z=1}^k \pi_z \mathcal{N}(x_i | \mu_z, \Sigma_z). \end{aligned}$$

- As usual, we'll take our objective function to be the log of this:

$$J(\pi, \mu, \Sigma) = \sum_{i=1}^n \log \left\{ \sum_{z=1}^k \pi_z \mathcal{N}(x_i | \mu_z, \Sigma_z) \right\}$$

Properties of the GMM Log-Likelihood

- GMM log-likelihood:

$$J(\pi, \mu, \Sigma) = \sum_{i=1}^n \log \left\{ \sum_{z=1}^k \pi_z \mathcal{N}(x_i | \mu_z, \Sigma_z) \right\}$$

- Let's compare to the log-likelihood for a single Gaussian:

$$\begin{aligned} & \sum_{i=1}^n \log \mathcal{N}(x_i | \mu, \Sigma) \\ &= -\frac{nd}{2} \log(2\pi) - \frac{n}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^n (x_i - \mu)' \Sigma^{-1} (x_i - \mu) \end{aligned}$$

- For a single Gaussian, the log cancels the exp in the Gaussian density.
 - \implies Things simplify a lot.
- For the GMM, the sum inside the log prevents this cancellation.
 - \implies Expression more complicated. No closed form expression for MLE.

Identifiability Issues for GMM

- Suppose we have found parameters

$$\text{Cluster probabilities : } \pi = (\pi_1, \dots, \pi_k)$$

$$\text{Cluster means : } \mu = (\mu_1, \dots, \mu_k)$$

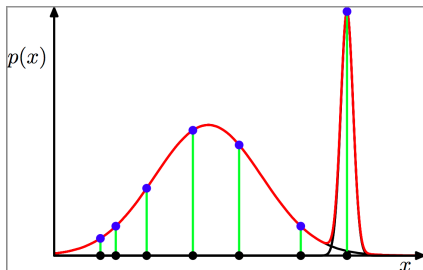
$$\text{Cluster covariance matrices: } \Sigma = (\Sigma_1, \dots, \Sigma_k)$$

that are at a local minimum.

- What happens if we shuffle the clusters? e.g. Switch the labels for clusters 1 and 2.
- We'll get the same likelihood. How many such equivalent settings are there?
- Assuming all clusters are distinct, there are $k!$ equivalent solutions.
- Not a problem per se, but something to be aware of.

Singularities for GMM

- Consider the following GMM for 7 data points:



- Let σ^2 be the variance of the skinny component.
- What happens to the likelihood as $\sigma^2 \rightarrow 0$?
- In practice, we end up in local minima that do not have this problem.
 - Or keep restarting optimization until we do.
- Bayesian approach or regularization will also solve the problem.

From Bishop's *Pattern recognition and machine learning*, Figure 9.7.

Gradient Descent / SGD for GMM

- What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = - \sum_{i=1}^n \log \left\{ \sum_{z=1}^k \pi_z \mathcal{N}(x_i | \mu_z, \Sigma_z) \right\}?$$

- Can be done – but need to be clever about it.
- Each matrix $\Sigma_1, \dots, \Sigma_k$ has to be positive semidefinite.
- How to maintain that constraint?
 - Rewrite $\Sigma_i = M_i M_i^T$, where M_i is an unconstrained matrix.
 - Then Σ_i is positive semidefinite.
- But we actually prefer positive definite, to avoid singularities.

Cholesky Decomposition for SPD Matrices

Theorem

Every symmetric positive definite matrix $A \in \mathbf{R}^{d \times d}$ has a unique **Cholesky decomposition**:

$$A = LL^T,$$

where L a **lower triangular matrix** with positive diagonal elements.

- A lower triangular matrix has half the number of parameters.
- Symmetric positive definite is better because avoids singularities.
- Requires a non-negativity constraint on diagonal elements.
 - e.g. Use projected SGD method like we did for the Lasso.

MLE for Gaussian Model

- Let's start by considering the MLE for the Gaussian model.
- For data $\mathcal{D} = \{x_1, \dots, x_n\}$, the log likelihood is given by

$$\sum_{i=1}^n \log \mathcal{N}(x_i | \mu, \Sigma) = -\frac{nd}{2} \log(2\pi) - \frac{n}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^n (x_i - \mu)' \Sigma^{-1} (x_i - \mu).$$

- With some calculus, we find that the MLE parameters are

$$\mu_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\Sigma_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^n (x_i - \mu_{\text{MLE}})(x_i - \mu_{\text{MLE}})^T$$

- For GMM, If we knew the cluster assignment z_i for each x_i ,
 - we could compute the MLEs for each cluster.

Cluster Responsibilities: Some New Notation

- Denote the probability that observed value x_i comes from cluster j by

$$\gamma_i^j = \mathbb{P}(Z = j \mid X = x_i).$$

- The **responsibility** that cluster j takes for observation x_i .
- Computationally,

$$\begin{aligned} \gamma_i^j &= \mathbb{P}(Z = j \mid X = x_i). \\ &= p(Z = j, X = x_i) / p(x) \\ &= \frac{\pi_j \mathcal{N}(x_i \mid \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i \mid \mu_c, \Sigma_c)} \end{aligned}$$

- The vector $(\gamma_i^1, \dots, \gamma_i^k)$ is exactly the **soft assignment** for x_i .
- Let $n_c = \sum_{i=1}^n \gamma_i^c$ be the number of points “soft assigned” to cluster c .

EM Algorithm for GMM: Overview

- 1 Initialize parameters μ, Σ, π .
- 2 “E step”. Evaluate the responsibilities using current parameters:

$$\gamma_i^j = \frac{\pi_j \mathcal{N}(x_i | \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i | \mu_c, \Sigma_c)},$$

for $i = 1, \dots, n$ and $j = 1, \dots, k$.

- 3 “M step”. Re-estimate the parameters using responsibilities:

$$\mu_c^{\text{new}} = \frac{1}{n_c} \sum_{i=1}^n \gamma_i^c x_i$$

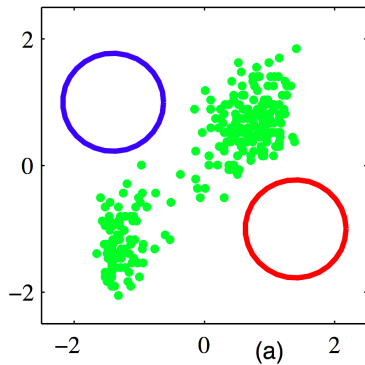
$$\Sigma_c^{\text{new}} = \frac{1}{n_c} \sum_{i=1}^n \gamma_i^c (x_i - \mu_{\text{MLE}}) (x_i - \mu_{\text{MLE}})^T$$

$$\pi_c^{\text{new}} = \frac{n_c}{n},$$

- 4 Repeat from Step 2, until log-likelihood converges.

EM for GMM

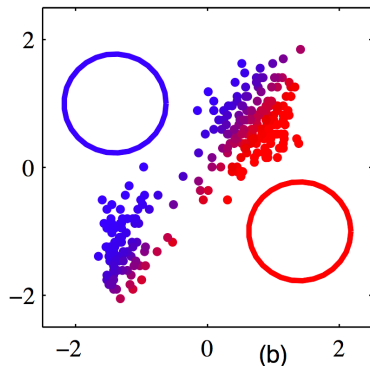
- Initialization



From Bishop's *Pattern recognition and machine learning*, Figure 9.8.

EM for GMM

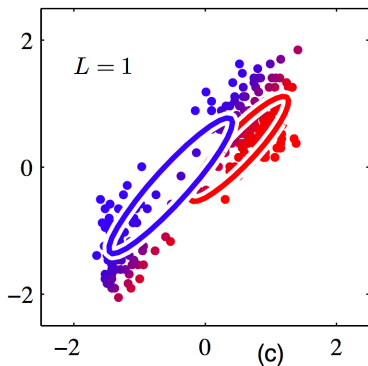
- First soft assignment:



From Bishop's *Pattern recognition and machine learning*, Figure 9.8.

EM for GMM

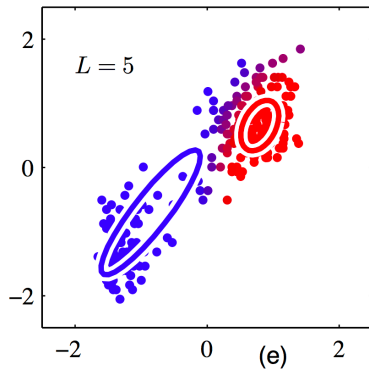
- First soft assignment:



From Bishop's *Pattern recognition and machine learning*, Figure 9.8.

EM for GMM

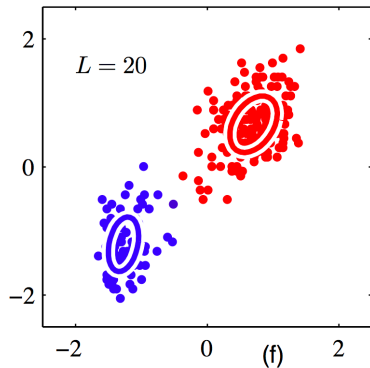
- After 5 rounds of EM:



From Bishop's *Pattern recognition and machine learning*, Figure 9.8.

EM for GMM

- After 20 rounds of EM:



From Bishop's *Pattern recognition and machine learning*, Figure 9.8.

Relation to K -Means

- EM for GMM seems a little like k -means.
- In fact, there is a precise correspondence.
- First, fix each cluster covariance matrix to be $\sigma^2 I$.
- As we take $\sigma^2 \rightarrow 0$, the update equations converge to doing k -means.
- If you do a quick experiment yourself, you'll find
 - Soft assignments converge to hard assignments.
 - Has to do with the tail behavior (exponential decay) of Gaussian.

Possible Topics for Next Time

- In last lecture, will give high level view of several topics.
- Possibilities:
 - General EM Algorithm.
 - Bandit problems.
 - LDA / Topic Models
 - Ranking problems.
 - Collaborative Filtering.
 - Generalization bounds.
 - Sequence models (maximum entropy Markov models, HMMs)